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FAILURE PREDICTION FOR AN ON-LINE MAINTENANCE SYSTEM IN A POISS--ETC(U)

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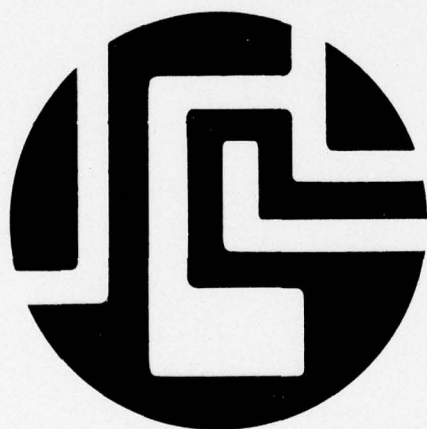
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FAILURE PREDICTION FOR AN ON-LINE MAINTENANCE SYSTEM  
IN A POISSON SHOCK ENVIRONMENT

by

K.-S. Lu

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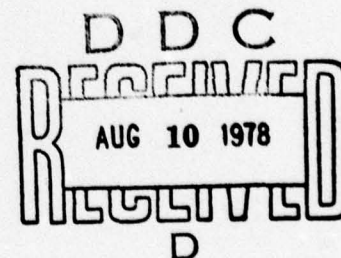
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Failure Prediction for an On-Line Maintenance Systems  
in a Poisson Shock Environment

Keh-Shew Lu\*

Department of Electrical Engineering  
Texas Tech University  
Lubbock, Texas 79409

January 1978



\*This research supported in part by ONR contract 75-C-0924. The author is now with the Texas Instruments Corp., Dallas, Texas.

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## ABSTRACT

An analog system subject to the Poisson Shock is modeled using past performance data. Failure Dynamics of the system is estimated by curve fitting techniques. Algorithms for fault prediction in an on-line maintenance process are described. Several sequential refinement schemes are introduced to improve fault prediction. Some formulas and properties of system's statistics have been developed. A decision rule is introduced which is based on the criteria of simultaneously maximizing lifetime and minimizing the cost of on-line failures. Poisson Shock generator is implemented by computer for simulation of the on-line maintenance process. The computer simulations of a perfect, no measurement errors and identical drifting parameters, system are presented. The simulations of an imperfect system are studied by adding a noise to the system performance data.



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## CHAPTER 1

### INTRODUCTION

Fault analysis processes have been and will continue to be very significant factors in the safety and reliability of electrical systems. This is especially true due to the following facts: a rapid advancement in the complexity and size of modern systems; increased availability and capabilities of computers; and rapidly changing technologies in integrated circuit fabrication. Due to this, fault analysis has become much more than an academic research topic. Fault analysis is applicable in an industrial environment to minimize cost, extend the lifetime of the overall system, control maintenance schedules, and effectively plan manpower needs.

The techniques of fault analysis (detection, location and prediction) have been developed independently in two different areas, digital and analog systems. In the last decade, considerable fault location research has been performed in the digital system area. Preparata, Metze and Chien (1) introduced an elegant graphical model and defined the properties of t-diagnosable systems. This model proved to be very powerful with the development of characterization theorems for the system by Hakimi and Amin (2, 3). In 1973, Russell and Kime (4) extended the Preparata-Metze-Chien



model into a more general case. However, such a generalization increases the complexity of the system.

Fault detection research for digital systems was introduced by Patterson and Metze (6) in 1972 and by Pradhan and Reddy (7) in 1973. Several different methods were developed for solving the problem of real-time fault detection in an asynchronous sequential circuit. After 1974, Maki and Sawin III (8, 9, 10) published a series of papers discussing the fault detection and fault-tolerant design for both synchronous and asynchronous sequential circuits.

The use of graphical models for dealing with analog circuits was first suggested by Kirchhoff in 1847 (11). Using Kirchhoff's model, Berkowitz (12) developed his concept of solvability for fault diagnosis in 1960. From the necessary conditions for solvability, a new concept, Key Subgraph (13), was introduced. Gayer (14) applied the Key Subgraph concept to fault isolation in simple linear solid state amplifiers.

Another aspect in the fault analysis process for an analog device is fault prediction. To accurately predict a fault, a device must be tested at periodic maintenance intervals. If the device fails or does not operate correctly, it is replaced immediately. The device may be assumed good if its characteristics are normal. However, if the characteristics are slightly out of tolerance, but

the device still operates correctly, one can attempt to predict if the device will fail before the next scheduled maintenance interval. If device failure is predicted, it can be replaced before failure occurs as part of planned preventive maintenance.

With the advent of the low-cost microprocessor, on-line fault prediction is possible and practical. A curve fitting algorithm for on-line fault prediction was first introduced by Saeks, Liberty and Tung (15, 16, 17) in 1975. It was assumed that prior life-time statistics for the system under test were known. Also, performance data of the system at each maintenance interval were collected. The application of these data to a second order polynomial equation resulted in an estimation of the time at which the component under test would exceed tolerance limits. Based on a criterion of simultaneously minimizing on-line failures and maximizing component lifetime, a decision as to whether or not the component should be replaced is made at each maintenance interval.

The disadvantages of this curve fitting algorithm are: the application is limited to failures due to permanent overstress, the second order polynomial is too simple to describe the performance of the component, and the prior lifetime statistics for the component are often not available.

Another area where an extensive research effort is being made is shock models and wear processes. Esary, Mar-

shall and Proschan (21, 22, 23) introduced a shock model for the life distribution of a component subjected to a sequence of shocks randomly occurring in the time according to a homogeneous Poisson process. They also considered the related shock models in which each shock caused a random amount of damage and failure occurred when the accumulated damage exceeded a specified threshold. This failure model is well known in modern reliability theory.

Employing the Poisson-Shock model, another curve fitting fault prediction algorithm which will overcome the disadvantages of the Saeks-Liberty-Tung algorithm will be discussed in the succeeding chapters. In Chapter 2, an analog system subjected to the Poisson Shock is modeled using past performance data of the analog system. Algorithms for fault prediction in an on-line maintenance process are described. Several sequential refinement schemes have been developed to improve fault prediction. In Chapter 3, a decision rule is developed which is based on the criterion of simultaneously maximizing lifetime and minimizing the cost of on-line failures. Simulations of the proposed on-line maintenance process are presented in Chapter 4.



## CHAPTER 2

### SYSTEM MODELING

#### 2.1 Failure Dynamics

The performance of an analog device subject to a series of discrete shocks (switching process, improper operation, etc...) may drift due to the shock damage. Let  $C(N)$  represent values of a particular component parameter, where the component time,  $N$ , denotes the number of shocks the component has received. In order to normalize the value of parameter  $C$ , a 1 or 0 is assigned to a "brand new" or "fail" stage respectively.

The normalized value of the parameter should have the following properties:

- (I)  $C(0) = 1$
- (II)  $C(\infty) = 0$  (2.1-1)
- (III) If  $L$  is the smallest integer that  $C(L) = 0$ , then  
 $C(N) = 0 \quad \forall N \geq L$ .

It is assumed that drifting parameters can be well described by a difference equation of the following form:

$$C(N + 1) = C(N) - f(N) \quad N < L \quad (2.1-2)$$

where  $f(N)$  is the particular failure dynamics.

Due to the simplicity of the calculation, it is practical to assume that  $f(N)$  has a polynomial form with order  $h$ . That is, eq. (2.1-2) can be rewritten as

$$\begin{aligned} C(N) - C(N+1) &= a_0 + a_1 N + \dots + a_h N^h \\ &= \sum_{i=0}^h a_i N^i \end{aligned} \quad (2.1-3)$$

where  $N < L$  and the  $\{a_i\}$  are constants and  $L$  is the smallest integer satisfying the condition

$$\sum_{j=0}^{L-1} \sum_{i=0}^h a_i j^i \geq 1 \quad (2.1-4)$$

Solving the difference eq. (2.1-3) with the boundary conditions (2.1-1), one obtains

$$C(N) = 1 - \sum_{j=0}^{N-1} \sum_{i=0}^h a_i j^i \quad \text{if } N < L \quad (2.1-5)$$

$$C(N) = 0 \quad \text{if } N \geq L$$

$L$  is termed the "life-time" of the component, that is, the number of shocks necessary to cause the component to fail.

Consider a simple example where  $f(N)$  is taken to be a first order polynomial; that is,

$$f(N) = C(N) - C(N+1) = a_0 + a_1 N \quad (2.1-6)$$

From equation (2.1-6) and boundary conditions (2.1-1),  $C(N)$  can be expressed as

$$C(N) = 1 - a_0 N - \frac{N(N-1)}{2} a_1 \quad \text{if } N < L$$

$$C(N) = 0 \quad \text{if } N \geq L$$
(2.1-7)

Then the life-time of this component is the smallest integer satisfying the equation

$$1 - a_0 L - \frac{L(L-1)}{2} a_1 \geq 0 \quad (2.1-8)$$

That is,

$$L \geq \frac{\sqrt{(2a_0 - a_1)^2 + 8a_1} - (2a_0 - a_1)}{2a_1} \quad (2.1-9)$$

Table (2.1-1) gives the life-time of the component, satisfying eq. (2.1-8), with different values of  $a_0$  and  $a_1$ . Fig. (2-1-1) shows the life-time curve as a function of  $a_0$  with different  $a_1$ 's.



Table (2.1-1)

Component's Lifetime with different values of  $a_0$  and  $a_1$ 

$L$ $a_0$	$a_1$	0.001	0.002	0.003	0.005	0.007	0.01	0.02	0.03	0.05	0.07	0.1
		>30	>30	>30	>30	>30	>30	>30	>30	>30	>30	>30
0.001		>30	>30	26	21	18	15	11	9	7	6	5
0.002		>30	>30	26	21	18	15	11	9	7	6	5
0.003		>30	>30	26	20	17	15	11	9	7	6	5
0.005		>30	30	25	20	17	15	11	9	7	6	5
0.007		>30	29	25	20	17	14	11	9	7	6	5
0.01		>30	28	24	19	17	14	11	9	7	6	5
0.02		30	24	21	17	15	13	10	8	7	6	5
0.03		25	21	19	16	14	12	10	8	7	6	5
0.05		18	16	15	13	12	11	9	8	6	6	5
0.07		14	13	12	11	10	10	8	7	6	5	5
0.1		10	10	9	9	9	8	7	6	6	5	5

$$L \geq \frac{\sqrt{(2a_0 - a_1)^2 + 8a_1 - (2a_0 - a_1)}}{2a_1}$$

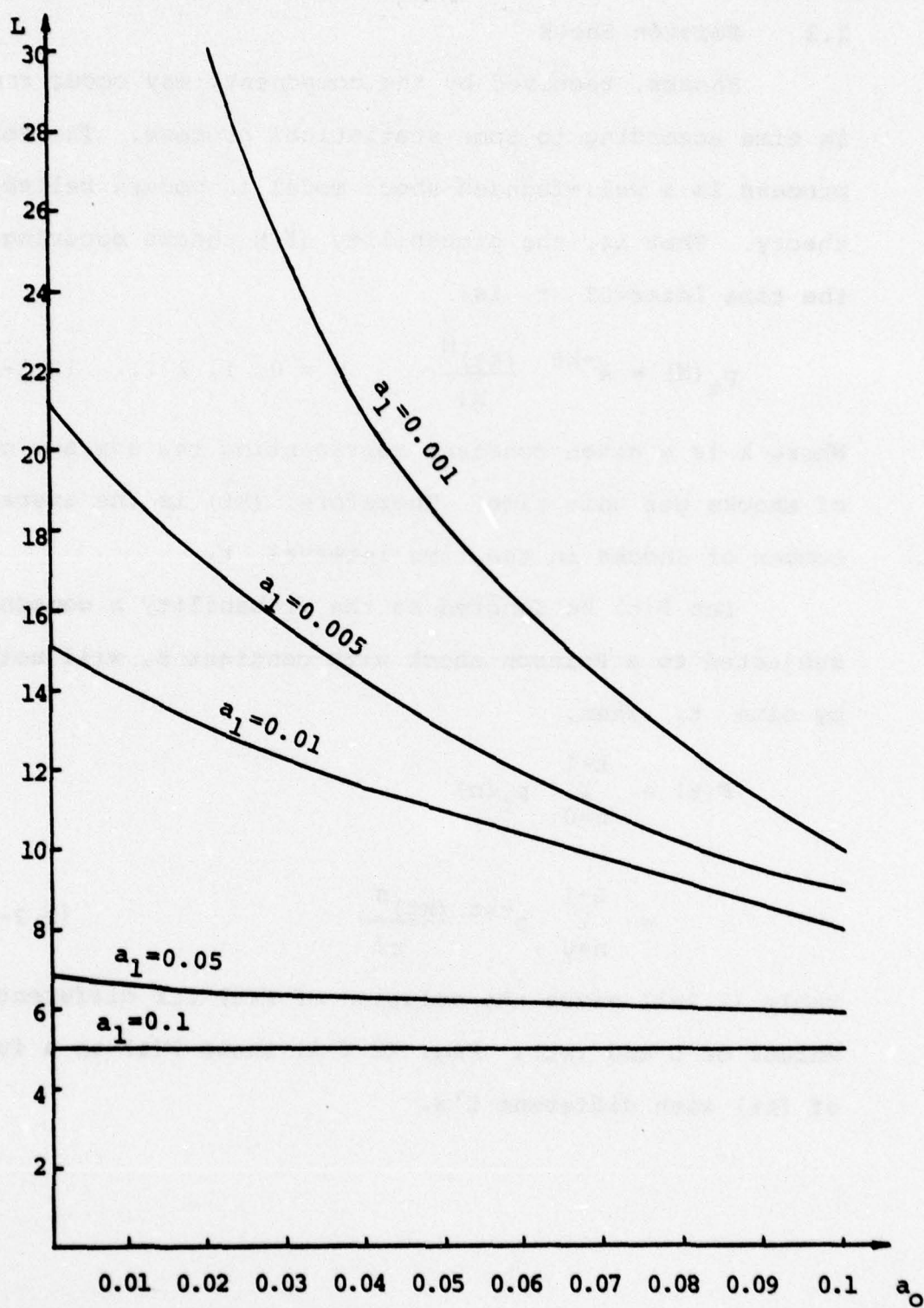


Fig. (2-1-1) Lifetime  $L$  v.s.  $a_0$  with different values of  $a_1$

## 2.2 Poisson Shock

Shocks, received by the component, may occur randomly in time according to some statistical process. The Poisson process is a well-founded shock model in modern reliability theory. That is, the probability of  $N$  shocks occurring in the time interval  $t$  is:

$$p_t(N) = e^{-kt} \frac{(kt)^N}{N!} \quad N = 0, 1, 2 \dots \quad (2.2-1)$$

Where  $k$  is a given constant representing the average number of shocks per unit time. Therefore,  $(kt)$  is the average number of shocks in the time interval  $t$ .

Let  $F(t)$  be denoted as the probability a component subjected to a Poisson shock with constant  $k$ , will not fail by time  $t$ . Then,

$$\begin{aligned} F(t) &= \sum_{n=0}^{L-1} p_t(n) \\ &= \sum_{n=0}^{L-1} e^{-kt} \frac{(kt)^n}{n!} \end{aligned} \quad (2.2-2)$$

Table (2.2-1) gives the solution of  $F(t)$  for different values of  $L$  and  $(kt)$ . Fig. (2-2-1) shows  $F(t)$  as a function of  $(kt)$  with different  $L$ 's.



TABLE (2.2-1)

Solutions of  $F(t)$  for different values of  $L$  and  $(kt)$ 

$F(t)\%$	$kt$										
	1	2	3	4	5	6	7	8	9	10	11
$L$	1	2	3	4	5	6	7	8	9	10	11
2	74	41	20	9	4	2	0.7	0.3	0.1	0	0
3	92	68	42	24	12	6	3	1	0.6	0.3	0.1
4	98	86	65	43	27	15	8	4	2	1	0.5
6	100	98	92	79	62	45	30	19	12	7	4
8	100	100	99	95	87	74	60	45	32	22	14
10	100	100	100	99	97	92	83	72	59	46	34
12	100	100	100	100	99	98	95	89	80	70	58
15	100	100	100	100	100	100	99	98	96	92	85
20	100	100	100	100	100	100	100	100	100	100	99
25	100	100	100	100	100	100	100	100	100	100	100
30	100	100	100	100	100	100	100	100	100	100	100

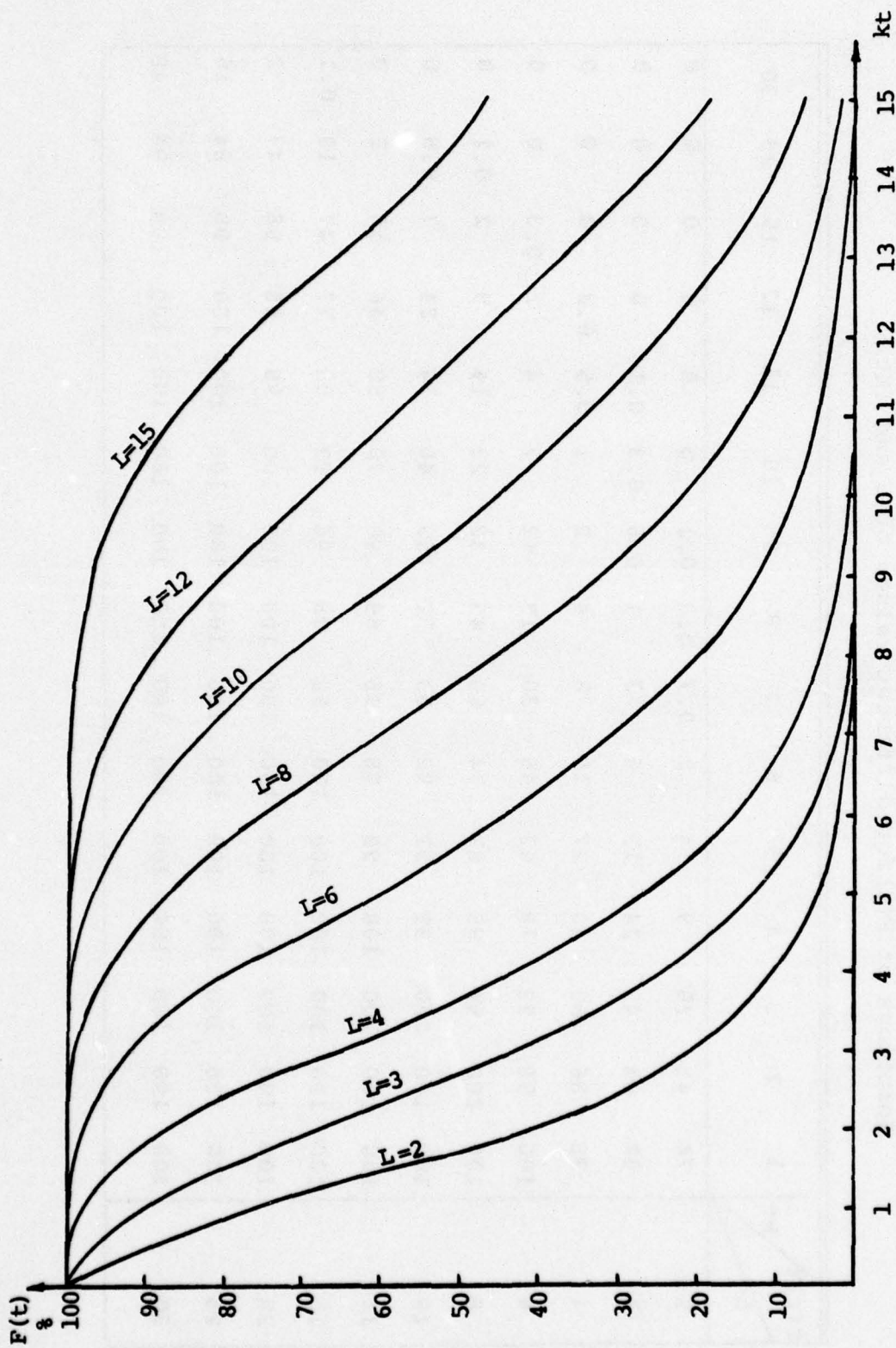


Fig. (2-2-1)  $F(t)$  v.s.  $(kt)$  with different lifetime

### 2.3 Estimation of the Failure Dynamics

In a periodic maintenance system, the performance data of a component is taken at each maintenance interval  $nT$ . This is,  $(C_1, C_2, \dots, C_g)$  is the performance data taken at maintenance times  $(T, 2T, \dots, gT)$ . The problem can be states as:

"Given performance data  $(C_1, C_2, \dots, C_g)$ , solve for the constants  $(a_0, a_1, \dots, a_h)$  of the failure dynamics in eq. (2.1-3)."

Since it is assumed that the system is subjected to Poisson Shock with constant  $k$ , that is the average number of shocks in each time interval  $T$  is  $kT$ . Therefore, eq. (2.1-5) can be rewritten as:

$$\sum_{j=0}^{mkT-1} a_0 j^0 + \sum_{j=0}^{mkT-1} a_1 j^1 + \dots + \sum_{j=0}^{mkT-1} a_h j^h = 1 - C_m \quad (2.3-1)$$

where  $m = 1, 2, 3, \dots, g$

or in matrix form:

$$\begin{bmatrix} \sum_{j=0}^{kT-1} j^0 & \sum_{j=0}^{kT-1} j^1 & \dots & \sum_{j=0}^{kT-1} j^h \\ \sum_{j=0}^{2kT-1} j^0 & \sum_{j=0}^{2kT-1} j^1 & \dots & \sum_{j=0}^{2kT-1} j^h \\ \dots & \dots & \dots & \dots \\ \sum_{j=0}^{gkT-1} j^0 & \sum_{j=0}^{gkT-1} j^1 & \dots & \sum_{j=0}^{gkT-1} j^h \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_h \end{bmatrix} = \begin{bmatrix} 1-C_1 \\ 1-C_2 \\ \vdots \\ 1-C_g \end{bmatrix} \quad (2.3-2)$$



or

$$J * A = Z \quad (2.3-3)$$

where

$$J = \begin{bmatrix} \sum_{j=0}^{kT-1} j^0 & \sum_{j=0}^{kT-1} j^1 & \dots & \sum_{j=0}^{kT-1} j^h \\ \sum_{j=0}^{2kT-1} j^0 & \sum_{j=0}^{2kT-1} j^1 & \dots & \sum_{j=0}^{2kT-1} j^h \\ \dots & \dots & \dots & \dots \\ \sum_{j=0}^{gkT-1} j^0 & \sum_{j=0}^{gkT-1} j^1 & \dots & \sum_{j=0}^{gkT-1} j^h \end{bmatrix}$$

$$Z = \begin{bmatrix} 1-C_1 \\ 1-C_2 \\ \vdots \\ 1-C_g \end{bmatrix} \quad \text{and} \quad A = \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_h \end{bmatrix}$$

The solution  $A$  can be discussed in two different cases:

(i) If  $J$  is nonsingular, then

$$A = J^{-1} * Z \quad (2.3-4)$$

(ii) If  $J$  is singular and full column rank, then

$$A = (J^T J)^{-1} J^T Z \quad (2.3-5)$$

## 2.4 Sequential Refinement Schemes

To approximate the failure dynamics by a polynomial, the proper choice of the order  $h$  is, in general, quite difficult and depends upon physical considerations and the engineering experience. Once  $h$  is preselected, employing the technique developed in the previous section, the polynomial coefficients to best approximate the failure dynamics can be calculated. The accuracy of the failure dynamics depends greatly on the choice of the order and can be improved by considering more measurements. To find a new set of coefficients for a different combination of  $h$  and  $g$ , the entire calculation procedure must be repeated from the very beginning. This repetition is impractical in the on-line maintenance system. Therefore, sequential refinement schemes for obtaining new sets of coefficients without having to repeat the entire calculation will be developed for the case in which  $g$  and  $h$  are allowed to vary. It is assumed that the coefficients matrix  $A$ , found from the last section for an  $h$  order polynomial with  $g$  measurements, are solutions for eq. (2.3-2). Three typical cases will be considered in the succeeding paragraphs.

Case 1:

Suppose that an additional measurement  $C_{g+1}$  is taken at the next scheduled maintenance interval  $(g+1)T$ , then eq. (2.3-2) becomes:

$$\begin{bmatrix}
 \sum_{j=0}^{kT-1} j^0 & \sum_{j=0}^{kT-1} j^1 & \dots & \sum_{j=0}^{kT-1} j^h \\
 \sum_{j=0}^{2kT-1} j^0 & \sum_{j=0}^{2kT-1} j^1 & \dots & \sum_{j=0}^{2kT-1} j^h \\
 \dots & \dots & \dots & \dots \\
 \sum_{j=0}^{gkT-1} j^0 & \sum_{j=0}^{gkT-1} j^1 & \dots & \sum_{j=0}^{gkT-1} j^h \\
 \hline
 \sum_{j=0}^{(g+1)kT-1} j^0 & \sum_{j=0}^{(g+1)kT-1} j^1 & \dots & \sum_{j=0}^{(g+1)kT-1} j^h
 \end{bmatrix}
 \begin{bmatrix}
 a_0 \\
 a_1 \\
 \dots \\
 a_h
 \end{bmatrix}
 =
 \begin{bmatrix}
 1-C_1 \\
 1-C_2 \\
 \dots \\
 1-C_g \\
 \hline
 1-C_{g+1}
 \end{bmatrix}
 \quad (2.4-1)$$

or

$$\begin{bmatrix}
 J \\
 \hline
 \hat{J}
 \end{bmatrix}
 * \hat{A} = \begin{bmatrix}
 Z \\
 \hline
 \hat{Z}
 \end{bmatrix} \quad (2.4-2)$$

where  $J$  and  $Z$  were defined in eq. (2.3-3)

$$\hat{J} = \begin{bmatrix}
 \sum_{j=0}^{(g+1)kT-1} j^0 & \sum_{j=0}^{(g+1)kT-1} j^1 & \dots & \sum_{j=0}^{(g+1)kT-1} j^h
 \end{bmatrix}$$

$$\hat{A} = [a_0 \ a_1 \ \dots \ a_h]$$

$$\hat{Z} = 1-C_{g+1}$$



The matrix  $\hat{A}$  can be solved as:

$$\hat{A} = (J^T J + \hat{J}^T \hat{J})^{-1} (J^T Z + \hat{J}^T \hat{Z}) \quad (2.4-3)$$

To avoid the calculation of a new matrix inverse as indicated in eq. (2.4-3) a sequential algorithm (24, 25) may be used to update  $\hat{A}$ . That is, assume

$$P = (J^T J)^{-1}$$

is given, then from eq. (2.3-5)

$$A = P J^T Z$$

and,

$$(J^T J + \hat{J}^T \hat{J})^{-1} = P - P \hat{J}^T (\hat{J} P \hat{J}^T + I)^{-1} \hat{J} P \quad (2.4-4)$$

Therefore,

$$\begin{aligned} \hat{A} &= P - P \hat{J}^T (\hat{J} P \hat{J}^T + I)^{-1} \hat{J} P (J^T Z + \hat{J}^T \hat{Z}) \\ &= A + P \hat{J}^T \{ I - (\hat{J} P \hat{J}^T + I)^{-1} \hat{J} P \hat{J}^T \} \hat{Z} - P \hat{J}^T (\hat{J} P \hat{J}^T + I)^{-1} \hat{J} A \end{aligned} \quad (2.4-5)$$

using the identity

$$I - (\hat{J} P \hat{J}^T + I)^{-1} \hat{J} P \hat{J}^T = (\hat{J} P \hat{J}^T + I)^{-1} \quad (2.4-6)$$

Eq. (2.4-5) becomes

$$\hat{A} = A + P \hat{J}^T (\hat{J} P \hat{J}^T + I)^{-1} (\hat{Z} - \hat{J} A) \quad (2.4-7)$$

Notice that the manipulation of the matrix  $(\hat{J} P \hat{J}^T + I)^{-1}$  is a simple scalar inversion.

Case 2:

Suppose an  $h$  order polynomial has been used and it is then determined that the error is too large. It would be more desirable to use a higher order,  $H$ , polynomial without having to repeat the entire calculation. Then, eq. (2.3-2) becomes:

$$\begin{bmatrix} \sum_{j=0}^{kT-1} j^0 & \sum_{j=0}^{kT-1} j^1 & \dots & \sum_{j=0}^{kT-1} j^h & \sum_{j=0}^{kT-1} j^H \\ \sum_{j=0}^{2kT-1} j^0 & \sum_{j=0}^{2kT-1} j^1 & \dots & \sum_{j=0}^{2kT-1} j^h & \sum_{j=0}^{2kT-1} j^H \\ \dots & \dots & \dots & \dots & \dots \\ \sum_{j=0}^{gkT-1} j^0 & \sum_{j=0}^{gkT-1} j^1 & \dots & \sum_{j=0}^{gkT-1} j^h & \sum_{j=0}^{gkT-1} j^H \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \dots \\ a_h \\ \dots \\ a_H \end{bmatrix} = \begin{bmatrix} 1-C_1 \\ 1-C_2 \\ \dots \\ 1-C_g \end{bmatrix} \quad (2.4-8)$$

or

$$\begin{bmatrix} J & \hat{J} \end{bmatrix} \begin{bmatrix} \hat{A} \\ -\frac{\hat{A}}{\hat{B}} \end{bmatrix} = Z \quad (2.4-9)$$

where  $J$  and  $Z$  were defined in eq. (2.3-3),

$$\hat{J} = \begin{bmatrix} \sum_{j=0}^{kT-1} j^{h+1} & \dots & \sum_{j=0}^{kT-1} j^H \\ \dots & \dots & \dots \\ \sum_{j=0}^{gkT-1} j^{h+1} & \dots & \sum_{j=0}^{gkT-1} j^H \end{bmatrix}$$

$$\hat{A} = \begin{bmatrix} a_0 \\ a_1 \\ \cdot \\ \cdot \\ a_h \end{bmatrix} \quad \text{and} \quad \hat{B} = \begin{bmatrix} a_{h+1} \\ a_{h+2} \\ \cdot \\ \cdot \\ a_H \end{bmatrix}$$

Assume A is the solution for the eq. (2.3-5), that is,

$$A = (J^T J)^{-1} J Z \quad (2.4-10)$$

The new solution matrices  $\hat{A}$ ,  $\hat{B}$  can be expressed as

$$\begin{aligned} \begin{bmatrix} \hat{A} \\ \hat{B} \end{bmatrix} &= \begin{pmatrix} \begin{bmatrix} J^T \\ \hat{J}^T \end{bmatrix} \\ \begin{bmatrix} J \\ \hat{J} \end{bmatrix} \end{pmatrix}^{-1} \begin{bmatrix} J^T \\ \hat{J}^T \end{bmatrix} Z \\ &= \begin{bmatrix} J^T J & J^T \hat{J} \\ \hat{J}^T J & \hat{J}^T \hat{J} \end{bmatrix}^{-1} \begin{bmatrix} J^T \\ \hat{J}^T \end{bmatrix} Z \end{aligned} \quad (2.4-11)$$



In order to improve the computational efficiency, an algorithm (24, 25) may be used to compute the matrix inverse. That is, assume

$$\begin{aligned} P &= (J^T J)^{-1} \\ E &= J^T \hat{J} \\ F &= \hat{J}^T \hat{J} \end{aligned} \quad (2.4-12)$$

$$\begin{aligned} \text{Then, } \begin{bmatrix} J^T & J \\ \hat{J}^T & \hat{J} \end{bmatrix}^{-1} &= \begin{bmatrix} P^{-1} & E \\ E^T & F \end{bmatrix}^{-1} \\ &= \begin{bmatrix} P(I + ESE^T P) & -PES \\ -SE^T P & S \end{bmatrix} \end{aligned} \quad (2.4-13)$$

where,

$$\begin{aligned} S &= (F - E^T P E)^{-1} \\ &= (\hat{J}^T \hat{J} - \hat{J}^T J P J^T \hat{J})^{-1} \end{aligned} \quad (2.4-14)$$

The new solution for  $\hat{B}$  is found as follows using eq. (2.4-11) and eq. (2.4-13)

$$\begin{aligned} \hat{B} &= -SE^T P J^T Z + S \hat{J}^T Z \\ &= S \hat{J}^T (Z - J P J^T Z) \\ &= S \hat{J}^T (Z - J A) \end{aligned} \quad (2.4-15)$$

The solution  $\hat{A}$  is found in a similar way, i.e.

$$\hat{A} = P(I + ESE^T P)J^T Z - PES\hat{J}^T Z$$

$$= PJ^T Z + PESE^T PJ^T Z - PES\hat{J}^T Z$$

(2.4-16)

$$= A - PE(S\hat{J}^T Z - SE^T PJ^T Z)$$

$$= A - PJ^T \hat{J} B$$

## Case 3:

Suppose the order of the polynomial  $h$  is chosen to equal  $g$ , and whenever an additional measurement is taken at next maintenance interval, the order of the polynomial will be increased by 1. Then, eq. (2.3-4) becomes:

$$\begin{bmatrix} \sum_{j=0}^{kT-1} j^0 & \sum_{j=0}^{kT-1} j^1 \dots & \sum_{j=0}^{kT-1} j^g & \sum_{j=0}^{kT-1} j^{g+1} \\ \sum_{j=0}^{2kT-1} j^0 & \sum_{j=0}^{2kT-1} j^1 \dots & \sum_{j=0}^{2kT-1} j^g & \sum_{j=0}^{2kT-1} j^{g+1} \\ \dots & \dots & \dots & \dots \\ \sum_{j=0}^{gkT-1} j^0 & \sum_{j=0}^{gkT-1} j^1 \dots & \sum_{j=0}^{gkT-1} j^g & \sum_{j=0}^{gkT-1} j^{g+1} \\ \hline \sum_{j=0}^{(g+1)kT-1} j^0 & \sum_{j=0}^{(g+1)kT-1} j^1 \dots & \sum_{j=0}^{(g+1)kT-1} j^g & \sum_{j=0}^{(g+1)kT-1} j^{g+1} \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \dots \\ a_g \\ \hline a_{g+1} \end{bmatrix} = \begin{bmatrix} 1-C_1 \\ 1-C_2 \\ \dots \\ 1-C_g \\ \hline 1-C_{g+1} \end{bmatrix}$$

or

(2.4-17)

$$\begin{bmatrix} J & | & H \\ \hline \hat{J} & | & G \end{bmatrix} \begin{bmatrix} \hat{A} \\ \hline \hat{B} \end{bmatrix} = \begin{bmatrix} Z \\ \hline \hat{Z} \end{bmatrix} \quad (2.4-18)$$

where  $J$ ,  $Z$  were defined in eq. (2.3-3),  $\hat{J}$ ,  $\hat{A}$ ,  $\hat{Z}$  were defined in eq. (2.4-2) with  $h = g$ , and,



$$H = \begin{bmatrix} \sum_{j=0}^{kT-1} j^{g+1} \\ \sum_{j=0}^{2kT-1} j^{g+1} \\ \dots \\ \sum_{j=0}^{gkT-1} j^{g+1} \end{bmatrix}$$

$$G = \sum_{j=0}^{(g+1)kT-1} j^{g+1} \quad \text{and} \quad \hat{B} = a_{g+1}$$

That is the matrices A and B can be expressed as

$$\begin{bmatrix} \hat{A} \\ \hat{B} \end{bmatrix} = \begin{bmatrix} J & H \\ \hat{J} & G \end{bmatrix}^{-1} \begin{bmatrix} Z \\ \hat{Z} \end{bmatrix} \quad (2.4-19)$$

A similar algorithm as used in case 2 can be employed (24, 25). That is,

$$\begin{bmatrix} J & H \\ \hat{J} & G \end{bmatrix}^{-1} = \begin{bmatrix} J^{-1}(I + HS\hat{J}J^{-1}) & -J^{-1}HS \\ -S\hat{J}J^{-1} & S \end{bmatrix} \quad (2.4-20)$$

where  $S = (G - \hat{J}J^{-1}H)^{-1}$

From eq. (2.4-19) and eq. (2.4-20), the new solution A, B can be solved as:

$$\begin{aligned}\hat{B} &= -\hat{S}\hat{J}^{-1}z + \hat{S}\hat{Z} \\ &= S(\hat{Z} - \hat{J}A)\end{aligned}\quad (2.4-21)$$

$$\begin{aligned}\hat{A} &= J^{-1}(I + HS\hat{J}^{-1})z - J^{-1}HS\hat{Z} \\ &= J^{-1}z + J^{-1}HS\hat{J}^{-1}z - J^{-1}HS\hat{Z} \\ &= A - J^{-1}H(\hat{S}\hat{Z} - \hat{S}\hat{J}^{-1}z) \\ &= A - J^{-1}H\hat{B}\end{aligned}\quad (2.4-22)$$

## CHAPTER 3

### FAULT PREDICTION

For the fault prediction algorithm it is assumed that the value of a component, drifting due to shock damages, is known at a fixed set of points in the "time domain". Using this data, one estimated the unknown failure dynamics for the parameter. This is then used to compute the component's lifetime in the "shock domain"; that is, the number of shocks required to cause a failure.

In the following paragraphs some formulas and properties of component's statistics will be developed. Using estimated lifetimes, some decision rules will be introduced to compute the optimal time at which to replace the component so as to minimize the cost function. Several terms are defined as follows:

- (1)  $L$  : The component's estimated lifetime in the shock domain; i.e. the number of shocks required to cause a failure.
- (2)  $T_r$  : The optimal time at which to replace the component so as to minimize the cost function.
- (3)  $P_f$  : The probability that the failure occurs before the replacement time  $T_r$ .
- (4)  $P_r$  : The probability that the component is replaced by a new component at replacement time  $T_r$ ; i.e. the probability that the component's lifetime is longer than  $T_r$ .



- (5)  $f_L(t)$ : The probability density function that the component receives the  $L^{\text{th}}$  shock at time  $t$ , where  $0 < t \leq T_r$ .
- (6)  $\tilde{T}_f$  : The expected lifetime of components which failed before replacement.
- (7)  $\hat{T}$  : The expected lifetime of components which either failed before  $T_r$  or were replaced at  $T_r$ .
- (8)  $T^*$  : The expected lifetime of components; i. e.  

$$T^* = T \Big| T_r \rightarrow \infty$$

Let  $p_i(t)$  represent the density function of the Poisson distribution with parameter  $(kt)$  and  $E_L(t)$  represent the corresponding distribution function; i.e.

$$p_i(t) = \frac{(kt)^i}{i!} e^{-kt} \quad i = 0, 1, 2, \dots \quad (3-1)$$

$$E_L(t) = \sum_{i=0}^{L-1} p_i(t) \quad (3-2)$$

Several properties can be stated as follows:

$$(\text{prop 1}) \quad P_r = E_L(T_r) \quad (3-3)$$

$$\begin{aligned} P_r &= \sum_{i=0}^{L-1} \frac{(kt)^i}{i!} e^{-kT_r} \\ &= \sum_{i=0}^{L-1} p_i(T_r) \\ &= E_L(T_r) \end{aligned}$$

$$(\text{prop 2}) \quad P_f = 1 - E_L(T_r) \quad (3-4)$$

$$\begin{aligned}
 P_f &= \sum_{i=L}^{\infty} P_i(T_r) \\
 &= 1 - \sum_{i=0}^{L-1} P_i(T_r) \\
 &= 1 - E_L(T_r)
 \end{aligned}$$

$$(\text{Prop 3}) \quad P_r + P_f = 1 \quad (3-5)$$

$$(\text{Prop 4}) \quad \int_0^{T_r} P_i(t) dt = \frac{1}{k} \{1 - E_{i+1}(T_r)\} \quad (3-6)$$

$$\begin{aligned}
 \int_0^{T_r} P_i(t) dt &= \int_0^{T_r} \frac{(kt)^i}{i!} e^{-kt} dt \\
 &= \frac{k^i}{i!} \int_0^{T_r} t^i e^{-kt} dt \quad (3-7)
 \end{aligned}$$

Using the identity eq. (3-8)

$$\int x^m e^{ax} dx = e^{ax} \sum_{r=0}^m (-1)^r \frac{m! x^{m-r}}{(m-r)! a^{r+1}} \quad (3-8)$$

Eq. (3-7) becomes

$$\begin{aligned}
 \int_0^{T_r} P_i(t) dt &= \frac{k^i}{i!} e^{-kt} \sum_{r=0}^i (-1)^r \frac{i! t^{i-r}}{(i-r)! (-k)^{r+1}} \bigg|_0^{T_r} \\
 &= \frac{k^i}{i!} \left\{ e^{-k \cdot 0} \frac{i!}{k^{i+1}} - e^{-kT_r} \sum_{r=0}^i \frac{i! T_r^{i-r}}{(i-r)! k^{r+1}} \right\} \\
 &= \frac{1}{k} \left\{ 1 - e^{-kT_r} \sum_{r=0}^i \frac{(kT_r)^{i-r}}{(i-r)!} \right\}
 \end{aligned}$$

$$= \frac{1}{k} \{1 - e^{-kT_r} \sum_{j=0}^i \frac{(kT_r)^j}{j!}\}$$

$$= \frac{1}{k} \{1 - \sum_{j=0}^i p_j(T_r)\}$$

$$= \frac{1}{k} \{1 - E_{i+1}(T_r)\}$$

$$(\text{Prop 5}) \quad f_L(t) = \frac{P_{L-1}(t)}{\frac{1}{k} (1 - E_L(T_r))} \quad (3-9)$$

where  $0 < t \leq T_r$

If the interval  $(0, T_r]$  is divided into  $N$  subintervals, it is seen that

$$\text{Prob} [(i-1)\Delta < t < i\Delta] = \frac{1}{N} \quad (3-10)$$

$$\text{where } \Delta = \frac{T_r}{N}$$

Assume  $\Delta$  is small enough such that the probability of more than one shock in the subinterval  $((i-1)\Delta, i\Delta]$  is equal to 0. The probability of one shock in the subinterval  $((i-1)\Delta, i\Delta]$  is equal to  $k\Delta$ , where  $k$  is the constant rate of Poisson shocks.

For a given  $t$  in the subinterval  $((i-1)\Delta, i\Delta]$ , the probability of the  $L^{\text{th}}$  shock in this subinterval is:



Prob  $\{L^{th}$  shock occurs in

$$((i-1)\Delta, i\Delta / ((i-1)\Delta < t \leq i\Delta)\}$$

$$= \text{Prob}\{(L-1) \text{ shocks occur in } (0, (i-1)\Delta)\}$$

$$* \text{Prob}\{\text{one shock occurs in } ((i-1)\Delta, i\Delta)\}$$

$$= \frac{[k(i-1)]^{L-1}}{(L-1)!} e^{-k(i-1)\Delta} * k\Delta$$

$$= P_{L-1}((i-1)\Delta) * k\Delta \quad (3-11)$$

From Baylar theory,

$$\text{Prob}\{\text{have } L^{th} \text{ in } ((i-1)\Delta, i\Delta)\}$$

$$= \frac{\text{Prob}\{(i-1)\Delta < t \leq i\Delta\} *}{\sum_{j=1}^N \text{Prob}\{(j-1)\Delta < \hat{t} \leq j\Delta\} *}$$

$$\frac{* \text{Prob}\{L^{th} \text{ shock occurs in } ((i-1)\Delta < t \leq i\Delta / ((i-1)\Delta < t \leq i\Delta)}{* \text{Prob}\{L^{th} \text{ shock occurs in } ((j-1)\Delta < \hat{t} \leq j\Delta / ((j-1)\Delta < \hat{t} \leq j\Delta)}$$

$$= \frac{\frac{1}{N} P_{L-1}((i-1)\Delta) k\Delta}{\sum_{j=1}^N \frac{1}{N} P_{L-1}((j-1)\Delta) k\Delta}$$

$$= \frac{p_{L-1}((i-1)\Delta)}{\sum_{j=1}^N p_{L-1}((j-1)\Delta)} \quad (3-12)$$

Since  $\Delta$  is very small, two approximations can be made as follows:

- (i)  $(i-1)\Delta \doteq t$
- (ii) Left hand side of eq. (3-12) is approached to

$$f_L(t) * \Delta$$

That is, eq. (3-12) can be rewritten as

$$f_L(t) = \frac{p_{L-1}(t)}{\sum_{j=1}^N p_{L-1}((j-1)\Delta)\Delta}$$

If  $\Delta \rightarrow 0$ , by definition of the integral the summation becomes an integral, such that

$$f_L(t) = \frac{p_{L-1}(t)}{\int_0^T p_{L-1}(\hat{t}) d\hat{t}} \quad (3-14)$$

From (Prop 4), eq. (3-14) becomes

$$f_L(t) = \frac{P_{L-1}(t)}{\frac{1}{k} \{1 - E_L(T_r)\}}$$

$$(Prop 6) \quad \hat{T}_f = \frac{L}{k} * \frac{1 - E_{L+1}(T_r)}{1 - E_L(T_r)}$$

Since  $\hat{T}_f$  is the expected lifetime of the components which failed before replacement,

$$\begin{aligned} \hat{T}_f &= \int_0^{T_r} t f_L(t) dt \\ &= \int_0^{T_r} \frac{t P_{L-1}(t)}{\frac{1}{k} \{1 - E_L(T_r)\}} dt \\ &= \frac{\int_0^{T_r} t * \frac{(kt)^{L-1}}{(L-1)!} e^{-kt} dt}{\frac{1}{k} \{1 - E_L(T_r)\}} \\ &= \frac{\frac{L}{k} \int_0^{T_r} \frac{(kt)^L}{L!} e^{-kt} dt}{\frac{1}{k} \{1 - E_L(T_r)\}} \\ &= \frac{L * \int_0^{T_r} P_L(t) dt}{\{1 - E_L(T_r)\}} \end{aligned} \quad (3-16)$$



From (Prop 4), eq. (3-16) becomes

$$\tilde{T}_f = \frac{L}{k} * \frac{\{1 - E_{L+1}(T_r)\}}{\{1 - E_L(T_r)\}}$$

$$(\text{Prop 7}) \quad \hat{T} = \frac{L}{k} \{1 - E_{L+1}(T_r)\} + T_r E_L(T_r) \quad (3-17)$$

$$\begin{aligned} \hat{T} &= P_f \tilde{T}_f + P_r T_r \\ &= \{1 - E_L(T_r)\} * \frac{L}{k} * \frac{1 - E_{L+1}(T_r)}{1 - E_L(T_r)} + T_r E_L(T_r) \\ &= \frac{L}{k} \{1 - E_{L+1}(T_r)\} + T_r E_L(T_r) \end{aligned}$$

$$(\text{Prop 8}) \quad T^* = \frac{L}{k} \quad (3-18)$$

$$E_i(T_r)]_{T_r=\infty} = \sum_{j=0}^{i-1} \frac{(kt)^j}{j!} e^{-kt} \Big|_{t=\infty} \div 0$$

i. e. The probability of having a finite number of shocks in an infinite time period is zero.

$$\begin{aligned} T^* = \hat{T}]_{T_r=\infty} &= \frac{L}{k} \{1 - E_{L+1}(T_r)\} + T_r E_L(T_r)]_{T_r=\infty} \\ &= \frac{L}{k} \end{aligned}$$

(Prop 9)

$$(1) \quad \frac{d(P_f)}{d(kT_r)} = p_{L-1}(T_r) \quad (3-19)$$

$$(2) \quad \frac{d(P_r)}{d(kT_r)} = -P_{L-1}(T_r) \quad (3-20)$$

$$P_r = E_L(T_r)$$

$$= \sum_{i=0}^{L-1} \frac{(kT_r)^i}{i!} e^{-kT_r}$$

$$= e^{-kT_r} + \sum_{i=1}^{L-1} \frac{(kT_r)^i}{i!} e^{-kT_r}$$

$$\therefore \frac{d(P_r)}{d(kT_r)} = -e^{-kT_r} + \sum_{i=1}^{L-1} \left\{ \frac{i(kT_r)^{i-1}}{i!} - \frac{(kT_r)^i}{i!} \right\} e^{-kT_r}$$

$$= \sum_{i=1}^{L-1} \frac{(kT_r)^{i-1}}{(i-1)!} e^{-kT_r} - \sum_{i=0}^{L-1} \frac{(kT_r)^i}{i!} e^{-kT_r}$$

$$= E_{L-1} - E_L$$

$$= -P_{L-1}(T_r)$$

$$\therefore P_f = 1 - P_r$$

$$\therefore \frac{d(P_f)}{d(kT_r)} = P_{L-1}(T_r)$$

$$(\text{Prop 10}) \quad \frac{d(kT_f)}{d(kT_r)} = \frac{\{1 - E_L(T_r)\} P_L(T_r) - \{1 - E_{L+1}(T_r)\} P_{L-1}(T_r)}{[1 - E_L(T_r)]^2}$$

(3-21)

$$k\tilde{T}_f = L * \frac{1 - E_{L+1}(T_r)}{1 - E_L(T_r)}$$

$$\frac{d(kT_f)}{d(kT_r)} = L * \frac{\{1 - E_L(T_r)\} p_L(T_r) - \{1 - E_{L+1}(T_r)\} p_{L-1}(T_r)}{\{1 - E_L(T_r)\}^2}$$

$$(\text{Prop 11}) \quad \frac{d(\hat{kT})}{d(kT_r)} = E_L(T_r) \quad (3-22)$$

$$\hat{T} = \frac{L}{k} \{1 - E_{L+1}(T_r)\} + T_r E_L(T_r)$$

$$k\hat{T} = L \{1 - E_{L+1}(T_r)\} + kT_r E_L(T_r)$$

$$\frac{d(k\hat{T})}{d(kT_r)} = L \{p_L(T_r)\} + kT_r (-p_{L-1}(T_r)) + E_L(T_r)$$

$$= L p_L(T_r) - kT_r p_{L-1}(T_r) + E_L(T_r)$$

Since,

$$L p_L(T_r) = L * \frac{(kT_r)^L}{L!} e^{-kT_r}$$

$$= (kT_r) * \frac{(kT_r)^{L-1}}{(L-1)!} e^{-kT_r}$$

$$= kT_r p_{L-1}(T_r)$$

$$\therefore \frac{d(k\hat{T})}{d(kT_r)} = E_L(T_r)$$



Once the estimated failure time and statistics of the component under study have been computed it remains to make a replacement decision. This should be based on the probability of the component's failure " $P_f$ ", the cost of replacing the component " $C_R$ ", and the cost of an on-line failure " $C_f$ ". That is, the cost function to be optimized can be expressed as either

$$\text{Cost} = \frac{1}{\hat{T}} (C_R + C_f P_f) \quad (3-23)$$

or

$$\text{Cost} = C_f P_f + C_R P_r \quad (3-24)$$

One possible replacement criterion is based on the cost of an on-line failure and average lifetime wastage due to replacing the component before its actual failure.

$$\text{Cost} = C_f P_f + C_W (kT^* - k\hat{T}) \quad (3-25)$$

where  $C_f$  and  $C_W$  are two weight constants. The first term of eq. (3-25) represents the cost of an on-line failure and the second term represents the lifetime wastage. The replacement time  $T_r$  which simultaneously minimizes the cost of an on-line failure and lifetime wastage should satisfy eq. (3-26) or eq. (3-27).

$$0 = C_f P_{L-1}(T_r) - C_W E_L(T_r) \quad (3-26)$$

therefore,

$$\left(\frac{C_f}{C_w} - 1\right) * \frac{(kT_r)^{L-1}}{(L-1)!} = \sum_{i=0}^{L-2} \frac{(kT_r)^i}{i!} \quad (3-27)$$

Table (3-1) gives the replacement time  $kT_r$  with different weight constants ( $C_f/C_w$ ) and different estimated lifetimes. Fig. (3-1) shows  $T_r$  curves being a function of  $L$  with different ( $C_f/C_w$ ).

Table (3-1)

The replacement time ( $kT_r$ ) for different ( $C_f/C_w$ ) and estimated lifetimes (L)

$C_f/C_w$	3	4	5	6	7	8	9	10	15	20
2.0	2.73	4.59	6.5	8.44	10.39	12.35	14.3	16.29	26.2	36.2
3.0	1.62	2.89	4.24	5.62	7.03	8.46	9.89	11.34	18.65	26.1
4.0	1.22	2.27	3.4	4.58	5.79	7.02	8.26	9.52	15.9	22.4
5.0	1.0	1.93	2.95	4.01	5.11	6.23	7.37	8.52	14.4	20.4
6.0	0.86	1.71	2.65	3.64	4.67	5.72	6.79	7.87	13.4	19.2
7.0	0.77	1.56	2.44	3.37	4.35	5.35	6.37	7.41	12.8	18.3
8.0	0.70	1.44	2.28	3.17	4.11	5.07	6.06	7.06	12.2	17.6
9.0	0.64	1.35	2.15	3.01	3.92	4.85	5.8	6.78	11.8	17.
10.0	0.60	1.27	2.05	2.88	3.76	4.67	5.6	6.55	11.5	16.8
15.0	0.46	1.03	1.71	2.46	3.25	4.08	4.93	5.8	10.4	15.1
20.0	0.38	0.9	1.52	2.22	2.96	3.74	4.54	5.37	9.7	14.3



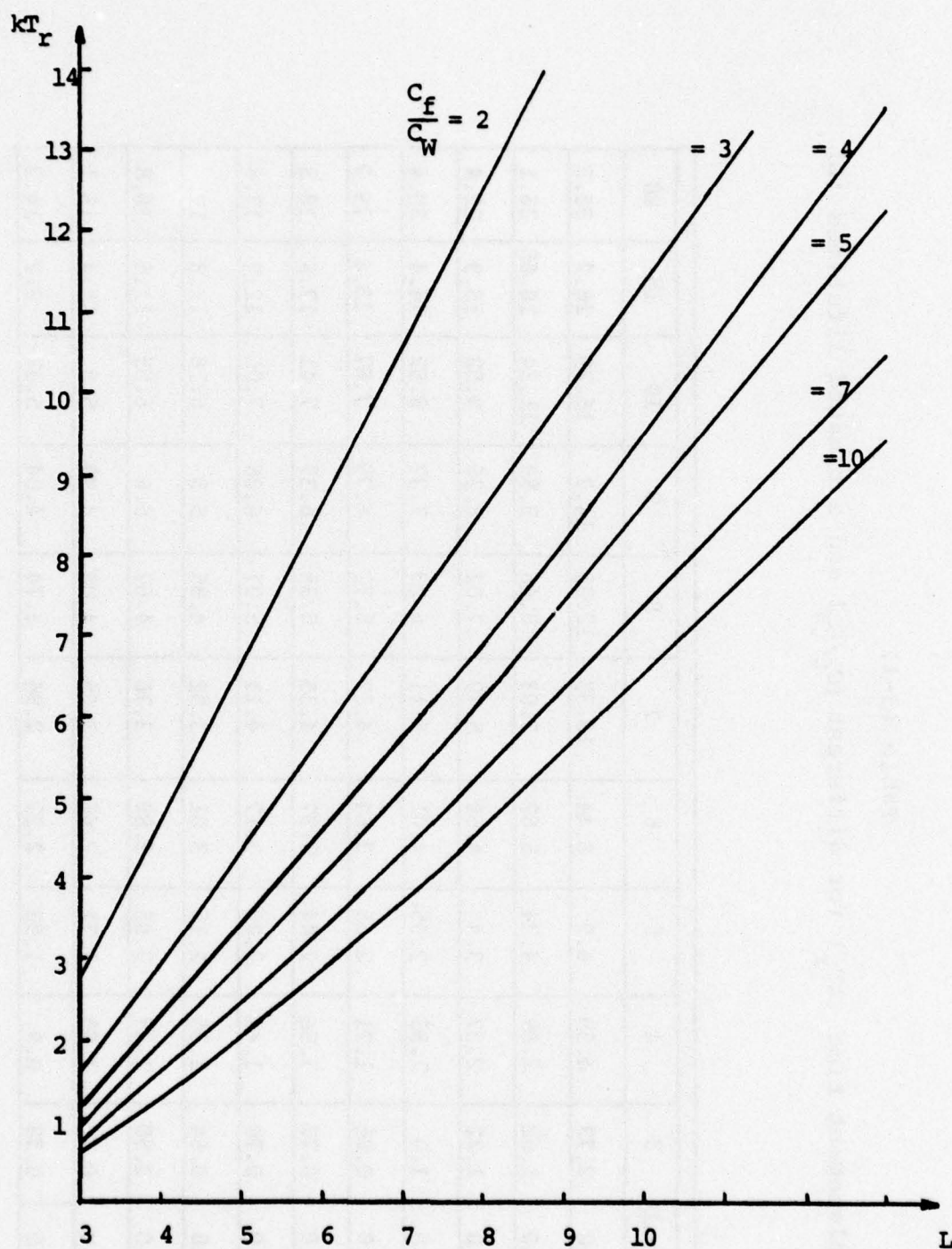


Fig. (3-1)

Replacement time ( $kT_r$ ) v.s. Lifetime  $L$  with different weight constant

## CHAPTER 4

### SIMULATIONS

#### 4.1 Poisson Shocks Generator

To generate a series of Poisson Shocks for the simulation of an on-line maintenance process by computer, uniformly distributed random numbers are needed. A linear congruential method (26, 27) can be used to generate numbers uniformly distributed on the interval (0, 1) by computer. That is, if the following relation exists between a set of numbers  $\{x_i\}$

$$x_{n+1} = (x_n * (2^s + 1) + b) \text{ mod } m \quad (4.1-1)$$

where  $s = Q/2$ ,  $m = 2^Q$

$Q$  = the number of shift register bits

in the computer

and  $b$  = any odd number

then, the sequence  $\{x_i\}$  is of period  $m$  and  $0 \leq x_i \leq m-1$ .

If each  $x_i$  is adjusted to the range (0, 1) by dividing by  $m$ ; i.e.,  $u_i = x_i/m$ , then  $u_i$  is a sequence of uniformly distributed random numbers between 0 and 1. That is,

$$p_U(u) = 1 \quad (4.1-2)$$

where  $0 \leq u < 1$

In order to convert the uniform distribution  $U$  to a Poisson distribution  $N$ , assume:

$$y = -\frac{1}{k} \ln(1-u) \quad (4.1-3)$$

or

$$u = 1 - e^{-ky} \quad (4.1-4)$$

where  $y \geq 0$

and  $k$  is a positive constant.

By the change of variable method (28, 29), the probability density function,  $p_Y(y)$ , becomes

$$\begin{aligned} p_Y(y) &= p_U(u) * \left| \frac{du}{dy} \right| \\ &= 1 * k e^{-ky} \end{aligned} \quad (4.1-5)$$

Let  $y$  represent the time interval between shocks, i.e., the times of successive shocks are  $y_1, y_2, \dots$  etc. Then, the distribution of the number of shocks,  $n$ , in a period  $T$  can be solved as follows:

$$\begin{aligned} \text{Prob}(N=n) &= \text{Prob}(y_1 + y_2 + \dots + y_{n+1} \geq T) \\ &\quad - \text{Prob}(y_1 + y_2 + \dots + y_n \geq T) \end{aligned} \quad (4.1-6)$$

From eq. (4.1-5) it can be seen that  $Y$  is distributed as  $\frac{1}{k} \chi^2$  with 2 degrees of freedom and  $\sum_{j=1}^m y_j$  is distributed as  $\frac{1}{k} \chi^2$  with  $2m$  degrees of freedom. Therefore, (30) eq. (4.1-6) becomes

$$\text{Prob}(N=n) = \text{Prob}(\chi^2_{2(n+1)} \geq 2kT) - \text{Prob}(\chi^2_{2n} \geq 2kT)$$



$$= \sum_{j=0}^n e^{-kT} \frac{(kT)^j}{j!} - \sum_{j=0}^{n-1} e^{-kT} \frac{(kT)^j}{j!}$$

(4.1-7)

$$= e^{-kT} * \frac{(kT)^n}{n!}$$

Hence N has a Poisson Distribution with parameter (kT).

Therefore, shocks that occur according to this Y distribution are called Poisson Shocks.

## 4.2 Simulation of a Perfect System

The computer simulation of a perfect (no measurement errors, and identical drifting parameters) on-line maintenance system can be performed using the following steps:

### Step 1 Define Constants

- (1)  $T$ : maintenance interval
- (2)  $k$ : time constant for Poisson Shock i.e. average number of shocks per unit time.
- (3)  $x_0$ : seed for uniform distributed random number generator
- (4)  $C_f/C_w$ : the ratio of two weight factors

### Step 2 Generate Poisson Shock

- (1) Generate Poisson Shock  $y_1, y_2, y_3, \dots$
- (2) Calculate number of shocks received in each maintenance interval  $n_1, n_2, n_3, \dots$

### Step 3 Define Values of Component Parameter

- (1)  $C(0) = 1$
- (2)  $C(i) = a_i$  where  $i$  is an integer and  $0 < i < L$  and  $0 < a_i \leq 1$
- (3)  $C(i) = 0$  where  $i$  is an integer and  $i \geq L$  --  $L$  is the lifetime of the component in the "Shock Domain".

Step 4 Observe Component Performance at Next  
Maintenance Time  $mT$

- (1) At maintenance time =  $mT$ , the total number of shocks received are  $N = \sum_{i=1}^m n_i$ , therefore the performance data

$$C_m = C(N)$$

- (2) Case 1: If  $C_m = 0$ , then an on-line failure occurs.

GO TO STEP 9

Case 2: If  $C_m > 0$ , then the component performs properly.

GO TO STEP 5

Step 5 Calculate Failure Dynamics

The new set of constants  $\{a_i\}$  of the failure dynamics can be calculated from eq. (2.4-23) and eq. (2.4-24)

Step 6 Solve Component's Lifetime

The lifetime of the component is then solved from eq. (2.1-5).

Step 7 Estimate Component's Replacement Time

The replacement time,  $T_r$ , can be predicted from eq. (3-27).

Step 8 Make a Decision

Case 1 : If  $T_r < (m+1)T$ , then the component will not be replaced.



GO TO STEP 4

Case 2: If  $T_r \geq (m+1)T$ , then the component will  
be replaced.

GO TO STEP 9

Step 9 Replace component

Replace the component and reset the count  
m to zero

GO TO STEP 4

**Example:**

The computer simulation of a on-line periodic maintenance system is performed for 600 maintenance intervals. Assume the system is subjected to a Poisson Shock with constant  $k=0.1$  shocks/hour and each interval  $T$  is 20 hours. The normalized values of a drifting parameter due to shocks damage are as followings:

$C(0) = 1.0$	$C(14) = 0.9$
$C(1) = 0.998$	$C(15) = 0.89$
$C(2) = 0.995$	$C(16) = 0.88$
$C(3) = 0.989$	$C(17) = 0.87$
$C(4) = 0.982$	$C(18) = 0.85$
$C(5) = 0.975$	$C(19) = 0.83$
$C(6) = 0.968$	$C(20) = 0.8$
$C(7) = 0.96$	$C(21) = 0.75$
$C(8) = 0.952$	$C(22) = 0.68$
$C(9) = 0.944$	$C(23) = 0.6$
$C(10) = 0.936$	$C(24) = 0.5$
$C(11) = 0.927$	$C(25) = 0.35$
$C(12) = 0.918$	$C(26) = 0.2$
$C(13) = 0.909$	$C(27) = 0.05$

and,  $C(N) = 0$  if  $N \geq 28$

Table (4.2-1) gives total number of replacements and failures within 600 maintenance intervals with different values of

$C_f/C_W$ . Table (4.2-2) shows the total number of replacements and failures with different replacement times.

Since the cost function is

$$\text{Cost} = C_f * P_f + C_W * (kT^* - k\hat{T})$$

the overall cost can be expressed as

$$\begin{aligned} \text{Cost} &= \frac{C_f}{C_W} * (\text{No. of failures}) \\ &+ 0.1 * (280 * (\text{No. of components used}) - 12000) \end{aligned}$$

The overall cost with different methods and different values of  $C_f/C_W$  are given in Table (4.2-3).



Table (4.2-1)

Total replacements and failures within 600  
maintenance intervals for different  $C_F/C_W$

$C_F/C_W$	No. of replacement	No. of failure
50	48	7
75	56	1
100	52	2
150	54	2
200	54	2

Table (4.2-2)

Total replacements and failures within 600 maintenance intervals for methods of constant time replacement

Constant replacement time	No. of replacement	No. of failure
every 6 intervals	100	0
every 7 intervals	85	0
every 8 intervals	75	0
every 9 intervals	65	1
every 10 intervals	59	1
every 11 intervals	48	6
every 12 intervals	39	11

Table (4.2-3)

Overall cost with different methods and different  $C_f/C_W$ 

Cost $\searrow$ Methods $\swarrow$ $C_f/C_W$	50	75	100	150	200
every 6 intervals	1600	1600	1600	1600	1600
every 7 intervals	1096	1096	1096	1096	1096
every 8 intervals	900	900	900	900	900
every 9 intervals	698	723	748	798	848
every 10 intervals	530	555	580	630	680
every 11 intervals	612	762	912	1212	1512
every 12 intervals	750	1025	1300	1850	2400
the algorithm	690	471	512	668	763



### 4.3 Simulation of an Imperfect System

The computer simulation of an imperfect on-line maintenance system can be performed by modifying the previous algorithm. Instead of using the same set of  $C(i)$  for every component, a noise is added into  $C(i)$  such that the lifetime of each component may not be the same but the average lifetime is still unchanged.

The total number of replacements and failures with  $C_f/C_W = 100$  in 600 maintenance intervals are solved by using different methods. The results with different noise levels are showed in Table (4.3-1).

Table (4.3-2) gives the overall cost with different methods and different noise levels.

Table (4.3-1)  
Overall cost with different noise levels

noise level method	20%		30%		40%		60%	
	No. of replace	No. of fail	No. of replace	No. of fail	No. of replace	No. of fail	No. of replace	No. of fail
every 6 intervals	100	0	100	0	100	0	94	6
every 7 intervals	85	0	85	0	84	1	78	8
every 8 intervals	75	0	72	3	71	4	64	12
every 9 intervals	64	2	63	3	60	7	52	17
every 10 intervals	56	4	51	9	45	15	45	18
every 11 intervals	45	10	45	10	45	10	39	20
every 12 intervals	36	15	35	16	36	17	31	23
the algorithm	56	3	55	5	55	5	50	11

Table (4.3-2)

Overall cost for different methods at different noise levels

cost noise level method	0 %	20 %	30 %	40 %	60 %
every 6 intervals	1600	1600	1600	1600	2200
every 7 intervals	1096	1096	1096	1280	2008
every 8 intervals	900	900	1200	1300	2128
every 9 intervals	748	848	948	1376	2432
every 10 intervals	580	880	1380	1980	2364
every 11 intervals	912	1340	1340	1340	2452
every 12 intervals	1300	1728	1828	1984	2612
the algorithm	512	752	980	752	1608



## CHAPTER 5

### CONCLUSION

In the preceeding chapters we have described a fault prediction algorithm for on-line maintenance systems. This algorithm essentially can be separated into three major steps: 1) applying a curve fitting technique to solve the failure dynamics by using past performance data, 2) predicting the lifetime of the system in the shock domain from the failure dynamics, 3) estimating the replacement time which simultaneously minimizes the cost of an on-line failure and lifetime wastage. In addition, sequential refinement schemes have been developed to solve the problem of inverting a potentially high dimensional matrix. Thus without having to repeat the entire calculation, the new sets of failure dynamics can be obtained recursively based on the old estimates and new data.

The algorithm has been tested in a variety of situations such as: perfect and imperfect system, different levels of noise, different sets of Poisson shocks. The results have been found to be surprisingly effective in predicting failures with relatively little wastage of lifetime and on-line failure cost.

Finally, similar algorithms for the replacement criterion based on the cost functions of eq. (4.1-23) and eq. (4.1-24) have been studied using the properties introduced in chapter 3. These algorithms yielded very good results.

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**APPENDIX**  
**FORTRAN PROGRAMS FOR THE SIMULATION SYSTEM**



// FOR MAX 01 JUN 77 23.409 HRS

\*LIST ALL

\*IOCS(CARD,DISK,1443 PRINTER)

\*ONEWORD INTEGERS

\*NONPROCESS PROGRAM

SUBROUTINE MAX(ICODE,AM,BM,CM,M,N,K)

C.

C. THIS IS THE SUBROUTINE FOR MATRIX OPERATION

C. AM(M,N) BM(N,K) CM(M,K)

C. ICODE=1 CM=AM+BM

C. ICODE=2 CM=AM-BM

C. ICODE=3 CM=AM\*BM

C.

DIMENSION AM(20,20),BM(20,20),CM(20,20)

IF(ICODE-2) 100,200,300

100 CONTINUE

DO 10 I=1,N

DO 10 J=1,N

CM(I,J) = AM(I,J) + BM(I,J)

10 CONTINUE

GO TO 400

200 CONTINUE

DO 20 I=1,N

DO 20 J=1,N

CM(I,J) = AM(I,J) - BM(I,J)

20 CONTINUE

GO TO 400

300 CONTINUE

DO 30 I=1,M

DO 30 J=1,K

CM(I,J) = 0.0

DO 5 L=1,N

CM(I,J) = CM(I,J) + AM(I,L)\*BM(L,J)

5 CONTINUE

30 CONTINUE

400 CONTINUE

RETURN

END

// FOR RMT 01 JUN 77 23.412 HRS

\*IOCS(CARD,DISK,1443 PRINTER)

\*LIST ALL

\*NONPROCESS PROGRAM

\*ONEWORD INTEGERS

INTEGER CR,PR

DIMENSION AM(20,20),BM(20,20),CM(20,20)

DIMENSION TSHUK(3000),CH(40)

DIMENSION E(1,20),H(20,1),A(20,1)

DIMENSION C(20,1),FIN(20,20)

DIMENSION DUM(20,20),DUM1(20,20)

DIMENSION DUM2(20,20),FINEW(20,20)

```

DIMENSION      NSK(600),IDESN(600), OB(20)
EQUIVALENCE(FIN(1,1),TSHUK(1) )
EQUIVALENCE(FIN(1,1),TSHUK(401),DUM2(1,1))
EQUIVALENCE ( DUM1(1,1),TSHUK(801) )
EQUIVALENCE(DUM(1,1),TSHUK(1201) )
EQUIVALENCE(A(1,1),TSHUK(1601))
EQUIVALENCE(H(1,1),TSHUK(1621))
EQUIVALENCE ( E(1,1),TSHUK(1641))
EQUIVALENCE(C(1,1),TSHUK(1661) )
EQUIVALENCE(OB( 1),TSHUK(1681))
EQUIVALENCE(CM(1,1),TSHUK(1701))
EQUIVALENCE(AM(1,1),TSHUK(2101))
EQUIVALENCE(BM(1,1),TSHUK(2501))
EQUIVALENCE(CH(1 ),TSHUK(2901))
CR=5
PR=6

```

```

C.
C. THIS SECTION IS TO GENERATED POISSON SHOCK
C.
C. SEED THE SEED FOR THE R. N. GENERATOR
  READ(CR,9101) SEED
9101  FORMAT(F 5.0)
C. NRND NUM OF RANDOM NUMBER TO BE GENERATED
  READ(CR,9101) NRND
9101  FORMAT(14)
C.
C. INPUT THE OBSERVATION PERIOD TPD
C.
  READ(CR,901) TPD
901  FORMAT(F15.5)
C. TCNST TIME CONSTANT
  READ (CR,9101) TCNST
  CNST=TPD/TCNST
  WRITE(PR,9106) SEED
9106  FORMAT(1H , 'THE SEED FOR THE RANDOM NO. IS ' ,F5.0 )
  WRITE(PR,9107) NRND
9107  FORMAT(1H , 'THE NO. OF POISSON SHOCK TO BE GENERATED IS ' , I5)
  WRITE(PR,9108) TCNST
9108  FORMAT(1H , 'AVE. NO. OF SHOCK PER OBSERVATION PERIOD IS ' ,F5.0)
C.
C.  $X(N+1) = (X(N) * AN + BN) \text{ MOD } CN$ 
C. X(N) IS R.N. N(0,16383)
C.
  DO 9115 INRND=1,NRND
  AN=129.0
  BN=111.0
  CN=16384.0
  S1=SEED*AN +BN
  IS=S1/CN
  S2=S1-IS*CN

```

```

      SEED=S2
C.
C. SEER NORMALIZED R. N.  N(0,1)
C. TSHUK IS POISSON SHUCK
C.
      SEEK=SEED/CN
      TSHUK(INRND) = (- CNST)*ALOG(1-SEER)
9115  CONTINUE
      JC= (NRND-1) /10 +1
      WRITE(PR,9130)
9130  FORMAT(1H , "THE SET OF POISSON SHOCK ARE ' )
      DO 9125 J=1, JC
      LL=(J-1)*10+1
      LH=J*10
      WRITE(PR,9135) (TSHUK(I), I=LL,LH)
9135  FORMAT (1H ,10(F10.5,1X))
9125  CONTINUE
C.
C. OUTPUT THE OBSERVATION PERIOD TPD
C.
      WRITE(PR,910) TPD
910   FORMAT(1H , "THE OBSERVATION PERIOD IS ' ,F15.5)
      I=0
      J=1
      LL=0
925   SUM=0.0
915   I=I+1
      IF(I-NRND) 918,918,930
918   SUM=TSHUK(I)+SUM
      IF(SUM-TPD) 915,920,920
920   IF(J-600) 924,924,922
922   WRITE(PR,923)
923   FORMAT(1H , 'THE DIMENSION OF NSK ARRAY IS TO SMALL')
      GO TO 930
924'  NSK(J) = I-LL-1
      LL=I-1
      TSHUK(I) = SUM-TPD
      I=I-1
      J=J+1
      GO TO 925
930   MAXPT=J-1
      DO 940 I=J,600
940   NSK(I)=0
      WRITE(PR,945)
945   FORMAT(1H , 'THE NO. OF SHOCK AT EACH OBSERVATION PERIOD ARE ' )
      JC= (MAXPT-1) /10 +1
      DO 950 J=1,JC
      LL=(J-1)*10 +1
      LH=J*10
      WRITE(PR,946) (NSK(I), I=LL,LH)

```



```

946  FORMAT(1H ,10 (I5,2X))
950  CONTINUE
      ICNP=5
      DO 100 IC=1,12
      ICNP=ICNP +1
      WRITE(PR,101) ICNP
101  FORMAT(//////,' PERIOD ',I4)
      ICNPT=0
      ISUM=0
      DO 105 I=1,MAXPT
      ICNPT=ICNPT +1
      ISUM=ISUM+NSK(I)
      IF(ISUM-28) 110,120,120
110  IF(ICNPT-ICNP) 105,111,111
111  WRITE(PR,115)I
115  FORMAT(' AT TEST POINT ', I 5,'REPLACE')
      ICNPT=0
      ISUM=0
      GO TO 105
120  WRITE(PR,125) I
125  FORMAT(' AT TEST POINT', I 5,' FAIL')
      ISUM=0
      ICNPT=0
105  CONTINUE
100  CONTINUE
C.
C.  INPUT THE CHARACTERISTIC OF THE DEVICE
C.
C.  C(0) = 1
C.  C(N+1) ALWAYS LESS THAN C(N)
C.  4 DATA CARDS WITH 10(F6.5 1X)
      WRITE(PR,9315)
9315  FORMAT(1H , "THE CHARACTERISTIC OF THE DEVICE ARE ')
      READ(CR,9310) I
9310  FORMAT(78X,A1)
      DO 9300 I=1,4
      LL=(I-1)*10 +1
      LH=I*10
      READ(CR,9350) (CH(J),J=LL,LH)
9350  FORMAT( 10(F6.5,1X))
      WRITE(PR,9360) (CH(J),J=LL,LH)
9360  FORMAT(1H ,10(F7.5,1X))
9300  CONTINUE
C.
C.  INPUT THE RATIO OF  CF/CW
C.  CODE FOR IDESN ARE
C.  IDESN(IDE) = 1  FAIL
C.  IDESN(IDE) = 2  CONTINUE
C.  IDESN(IDE) = 3  REPLACE

```

BLAN

C.  
 MT=TCNST  
 READ(CR,3002) CFW  
 3002 FORMAT(F15.5)  
 WRITE(PR,3003) CFW  
 3003 FORMAT(1H , 'THE RATIO CF/CW IS ' , F15.5)  
 MT=TCNST  
 IDE=0  
 3005 ICT=0  
 3006 IDE=IDE+1  
 IF (IDE-MAXPT) 3010,3010,4000  
 3010 ICT=ICT+1  
 IF (ICT-1) 3015,3015,3115  
 3015 ITSK= NSK(IDE)  
 IF (CH(ITSK)) 3020,3020,3025

C.  
 C. FIRST TIME TEST AND FAIL

C.  
 3020 WRITE(PR,3022) IDE,ITSK  
 3022 FORMAT(1H , "AT TEST POINT",I5,3X,  
 1' THE DEVICE TOTAL RECEIVED',I5,2X,  
 2' SHOCKS AND FAIL TO OPERATE')  
 IDESN(IDE)=1  
 GO TO 3005

C.  
 C. FIRST TIME BUT NOT FAIL

C.  
 3025 FIN(1,1)=1.0/MT  
 OB(1,1) = CH(ITSK)  
 C(1,1) = 1.0 -OB(1,1)  
 NORED=1  
 A(1,1)=FIN(1,1)\*C(1,1)  
 GO TO 502

C.  
 C. THIS IS THE ROUTINE TO SOLVE THE PARAMETERS

C NORED THE ORDER OF THE COEFFICIENTS

C. FI THE INVERSE OF THE MATRIX F

C. F X A = C

C. FH \* A = C

C. EG \* B = D

C. H (NORED , 1)

C. E (1 , NORED)

C. G , D IS CONSTANT

3115 CONTINUE  
 ITSK=ITSK + NSK(IDE)  
 IF (CH(ITSK)) 3020,3020,1000

1000 CONTINUE

C.  
 C. CALCULATE H C



```

C.      NOREL=NORED+1
        SUM=0.0
        DO 1010 I=1,NORED
          LML=(I-1) * MT
          IF (LML) 1002,1002,1003
1002     LML=1
1003     LMH= I*MT - 1
          DO 1005 J= LML , LMH
            SUM = SUM + J ** NOREL
1005     CONTINUE
          H(I,1)=SUM
          C(I,1)=1.0- OB(I)
1010     CONTINUE
C.
C.      CALCULATE G
C.
        LML= LMH+1
        LMH= LML+ MT
        DO 1110 J = LML ,LMH
          SUM= SUM + J ** NOREL
1110     CONTINUE
        G=SUM
C.
C.      CALCULATE E
C.
        DO 1111 I=2,NORED
1111     E(1,I) = 0.0
          E(1,1) = (NOREL*MT)
          LMH= (NOREL ) * MT - 1
          DO 1120 J= 2,NORED
            DO 1115 I= 1, LMH
              E(1, J) = E(1,J) + I ** J
1115     CONTINUE
1120     CONTINUE
          OB(NOREL) = CH(ITSK)
          D=1.0-OB(NOREL)
          CALL MAX (3,E,FIN,DUM,1,NORED,NORED)
          CALL MAX (3,DUM,H,DUM1,1,NORED,1)
          S=1.0/(G-DUM1(1,1))
          FINEW(NOREL,NOREL) =S
          DO 1500 J=1,NORED
            FINEW(NOREL,J) = -S *DUM(1,J)
1500     CONTINUE
          DO 1510 J=1,NORED
            DUM(1,J) = DUM(1,J) *S
1510     CONTINUE
          CALL MAX ( 3,FIN,H,DUM1,NORED,NORED,1)
          CALL MAX(3,DUM1,DUM,DUM2,NORED,1,NORED)
          CALL MAX (1,FIN,DUM2,FINEW,NORED,NORED,NORED)
          CALL MAX (3,FIN,H,DUM,NORED,NORED,1)

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DO 1520 I=1,NORED
FINEW(I,NORED) = -DUM(I,1) *S
1520 CONTINUE
C(NORED,1) =1.0-OB(NORED)
CALL MAX(3,FINEW ,C,A,NORED,NORED,1)
NORED=NORED
DO 400 I=1,NORED
DO 400 J=1,NORED
FIN(I,J) = FINEW(I,J)
400 CONTINUE
C.
C. FROM THE SET OF COEFFICIENTS A(I) SOLVE THE LIFE TIME
C.
502 SUM=0.0
WRITE(PR,7918) (A(IT,1) ,IT=1,NORED)
7918 FORMAT(' ***** A **',6(F9.4,1X))
J=0
SUM=SUM+A(1 ,1)
IF(SUM-1.0) 505,510,510
501 SUM=SUM+A(1,1)
IF( NORED-1) 500,500,513
513 DO 500 I=2,NORED
RJ=J
SUM1=RJ**(I-1)
SUM=SUM+A(I,1)*J**(I-1)
500 CONTINUE
IF(SUM-1.0) 505,510,510
505 J=J+1
IF(J-60) 501,501,510
510 LIFE=J+1
LIFEN=LIFE-ITSK
TLIFE=LIFEN/TCNST
WRITE(PR,7900)LIFEN
7900 FORMAT(' ** ** ** LIFE= ',15, 'SHOCK')
C.
C. USING NEWTON'S METHOD TO SOLVE TR
C. CFW CF/CW
C.
IF(LIFE -2) 550,560,565
550 TR=0.0
GO TO 700
560 TR=1.0/(CFW-1.0)
GO TO 700
565 IF(LIFE-3) 570,570,580
570 TR=(1.0 + SQRT(2.0*CFW-1))/(CFW-1)
GO TO 700
580 IF(LIFE-60) 590,567,567
567 TR=9999.0
GO TO 700
590 TR=LIFE*2.0

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      ICODP=0
      LIF1=LIFE-1
      LIF2=LIFE-2
      FP=1.0
594  FAT=1.0
      SUM=1.0
      TERM=1.0
      DO 591 I=1,LIF1
      TERM=TERM*TR/I
      SUM=SUM+TERM
591  CONTINUE
      F=CFW*TERM-SUM
      IF (F-0.0) 592,700,593
593  TR=TR-1.0
      FP=F
      ICODP=1
      GO TO 594
592  TR=TR- (F/(FP-F))
      LIF1=LIFE-1
      LIF2=LIFE-2
610  FAT=1.0
      SUM=1.0
      TERM=1.0
      DO 600 I=1,LIF2
      TERM=TERM*TR/I
      SUM=SUM+TERM
600  CONTINUE
      DF=CFW*TERM-SUM
      TERM=TERM*TR/LIF1
      SUM=SUM+TERM
      F=CFW*TERM-SUM
      TRN=TR-F/DF
      IF (ABS (TRN-TR) -0.001) 700,700,690
690  TR=TRN
      GO TO 610
700  CONTINUE
      TR=TR*TPD/TCNST
      TR=TR-TPD*ICT
      WRITE (PR,710)  IDE,ITSK,OB(ICT)
710  FORMAT (LH ,/, ' AT TEST POINT',I5, 2X,
1'THE DEVICE TOTAL RECEIVED',I5,2X,
2'SHOCKS AND THE MEASUREMENT IS',F15.5)
      WRITE (PR,715)  TLIFE, TR
715  FORMAT (' THE ESTIMATED LIFE TIME IS ',F9.3,' THE REPLACE TIME IS
1 ',F7.2)
      IF (TR-TPD) 720,720,730
C.
C.  REPLACE DEVICE
C.
720  WRITE (PR,722)

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722   FORMAT(1H+,T75,'LATER.  WE WILL REPLACE IT NOW')  
      IDESN(IDE)=3  
      GO TO 3005

C.

C.  CONTINUE USING THE DEVICE

C.

730   WRITE (PR,732)

732   FORMAT(1H+,T75,'LATER.  WE WILL CONTI. USE IT')  
      IDESN(IDE)=2  
      GO TO 3006

4000   CONTINUE  
      CALL EXIT  
      END